

# Analysis and Implementation of RC4 Stream Cipher

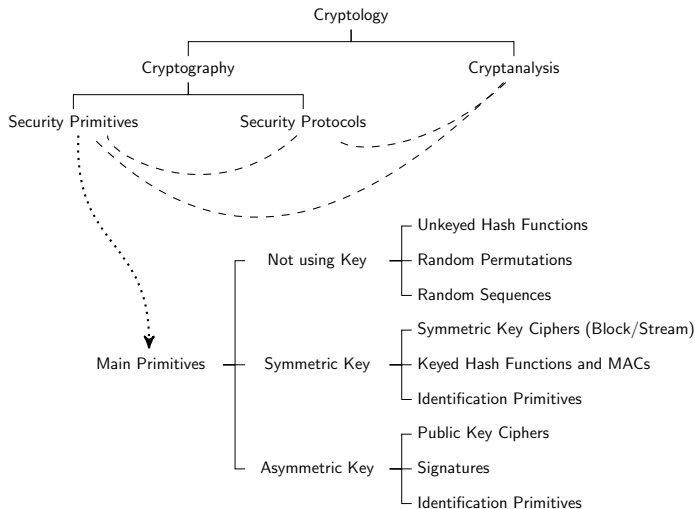
SOURAV SEN GUPTA

A thesis presented to Indian Statistical Institute in fulfillment of the thesis requirement for the degree of Doctor of Philosophy in Computer Science.

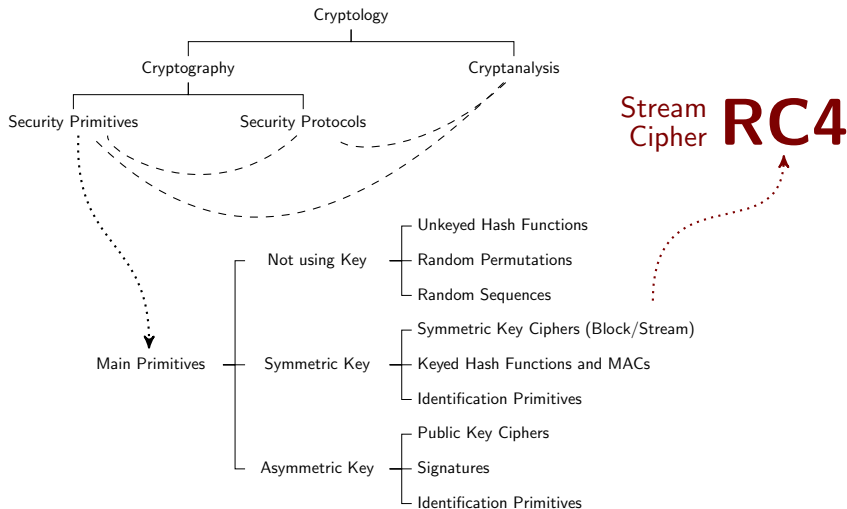
Under the supervision of  
Professor Subhamoy Maitra  
Applied Statistics Unit, ISI Kolkata

# Scope of the Thesis

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# Scope of the Thesis



## Organization of the Thesis

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There are 9 chapters, distributed over 2 major parts, in this thesis.

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### Chapter 1 – Preliminaries and Motivation

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Part I – Analysis of RC4

Chapter 2 – Overview of RC4 Analysis

Chapter 3 – Keylength biases

Chapter 4 – State-dependent biases

Chapter 5 – Keystream biases

Part II – Implementation of RC4

Chapter 6 – Overview of RC4 Designs

Chapter 7 – Design 1 (loop unrolling)

Chapter 8 – Design 2 (pipelining)

### Chapter 9 – Conclusion and Open Problems

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We deal with 10 research problems in this thesis.

We present 10 open problems in related research.



# Organization of this Talk

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## Introduction

- Stream Ciphers and RC4

- Motivation of our work

## Analysis of RC4 Stream Cipher

- Non-randomness in the output keystream

- Keylength dependent biases in RC4

- Long-term biases in RC4

- Biases related to the state-variables

- Contributions in RC4 Analysis

## Implementation of RC4 Stream Cipher

- Design based on loop unrolling

- Design based on hardware pipelining

- Contributions in RC4 Implementation

## Conclusion



# Stream Ciphers and RC4

## Stream Ciphers

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Exploit the notion of *perfect secrecy* by Shannon, 1949.

random keystream



plaintext message

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encrypted message

Encrypted message reveals no information about the plaintext for a *one-time-pad* encryption.

Shannon, Claude E. (October 1949). "Communication Theory of Secrecy Systems".  
Bell System Technical Journal (USA: AT&T Corporation) 28 (4):656–715.

## Stream Ciphers

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They aim at producing  
*long* random keystream  
from a *short* secret key.

plaintext message  
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encrypted message

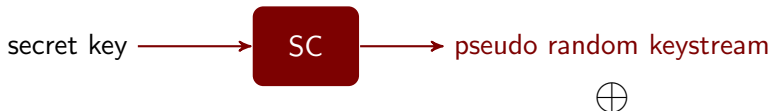
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Exploit the notion of *perfect secrecy* by Shannon, 1949.



They aim at producing  
*long* random keystream  
from a *short* secret key.

plaintext message

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encrypted message

**But never produce a truly random keystream!**

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## RC4 Stream Cipher

---

- Designed by Ron Rivest in 1987
- Description public in 1994 (?)

### POPULARITY

- Most used commercial cipher!
- Used in WEP, WPA, SSL/TLS.
- Numerous academic publications and patents.



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### POPULARITY

- Most used commercial cipher!
- Used in WEP, WPA, SSL/TLS.
- Numerous academic publications and patents.
- **Simplest cipher to describe!**



## RC4 Stream Cipher

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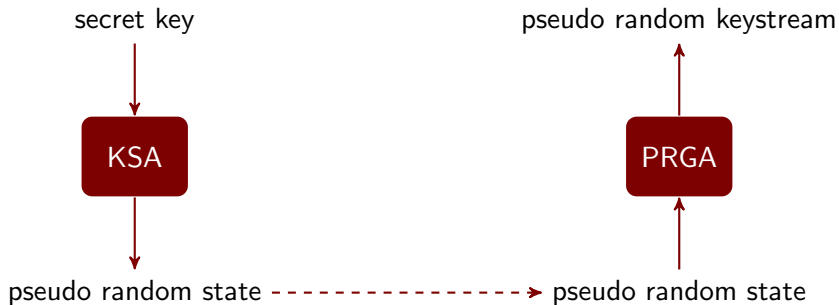


# RC4 Stream Cipher

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Key Scheduling Algorithm

Pseudo-Random Generation Algorithm



# RC4 Stream Cipher

---

## Key Scheduling Algorithm

256 bytes long  
secret key



pseudo random state  
permutation of  $\{0, 1, \dots, 255\}$

## Pseudo-Random Generation Algorithm

1 byte per iteration  
pseudo random keystream

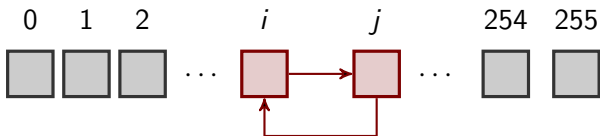


pseudo random state  
permutation of  $\{0, 1, \dots, 255\}$



## Key Scheduling Algorithm (KSA)

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Initialize index:  $j = 0$ ;

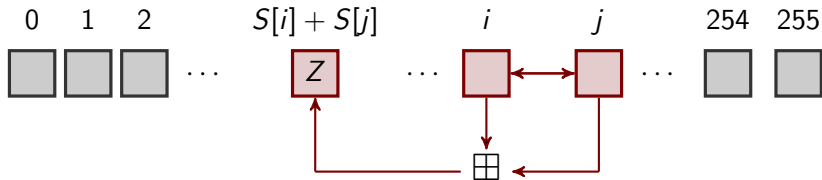
```

for  $i = 0, \dots, 255$  do
     $j = j + S[i] + K[i]$ ;
    Swap  $S[i] \leftrightarrow S[j]$ ;
end
  
```

INPUT:  $S$ -array initialized to identity permutation, and key  $K$

OUTPUT: Scrambled  $S$ -array

## Pseudo-Random Generation Algorithm (PRGA)



Initialize indices:  $i = j = 0$ ;

**while** *TRUE* **do**

$i = i + 1$ ;

$j = j + S[i]$ ;

Swap  $S[i] \leftrightarrow S[j]$ ;

Output  $Z = S[S[i] + S[j]]$ ;

**end**

INPUT: Scrambled  $S$ -array,  
obtained as the KSA output

OUTPUT: Pseudo-random stream



## RC4 toy example

---

KSA with  $N = 8$

$K = [3, 1, 5, 2, 7, 0, 6, 4]$

$S = [0, 1, 2, 3, 4, 5, 6, 7]$

$i = 0, j = 3$

$S = [3, 1, 2, 0, 4, 5, 6, 7]$

$i = 1, j = 5$

$S = [3, 5, 2, 0, 4, 1, 6, 7]$

$i = 2, j = 4$

$S = [3, 5, 4, 0, 2, 1, 6, 7]$

$i = 3, j = 6$

$S = [3, 5, 4, 6, 2, 1, 0, 7]$

$i = 4, j = 7$

$S = [3, 5, 4, 6, 7, 1, 0, 2]$

$i = 5, j = 0$

$S = [1, 5, 4, 6, 7, 3, 0, 2]$

$i = 6, j = 6$

$S = [1, 5, 4, 6, 7, 3, 0, 2]$

$i = 7, j = 4$

$S = [1, 5, 4, 6, 2, 3, 0, 7]$

## RC4 toy example

---

PRGA with  $N = 8$

K is no more required

$S = [1, 5, 4, 6, 2, 3, 0, 7]$

$i = 1$ , $j = 5$	$S = [1, 3, 4, 6, 2, 5, 0, 7]$	$Z = 1$
$i = 2$ , $j = 1$	$S = [1, 4, 3, 6, 2, 5, 0, 7]$	$Z = 7$
$i = 3$ , $j = 7$	$S = [1, 4, 3, 7, 2, 5, 0, 6]$	$Z = 5$
$i = 4$ , $j = 1$	$S = [1, 2, 3, 7, 4, 5, 0, 6]$	$Z = 0$
$i = 5$ , $j = 6$	$S = [1, 2, 3, 7, 4, 0, 5, 6]$	$Z = 0$
$i = 6$ , $j = 3$	$S = [1, 2, 3, 5, 4, 0, 7, 6]$	$Z = 4$
$i = 7$ , $j = 1$	$S = [1, 6, 3, 5, 4, 0, 7, 2]$	$Z = 1$
$i = 8$ , $j = \dots$	$S = \dots$	

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**such enigmatic a flair?!**

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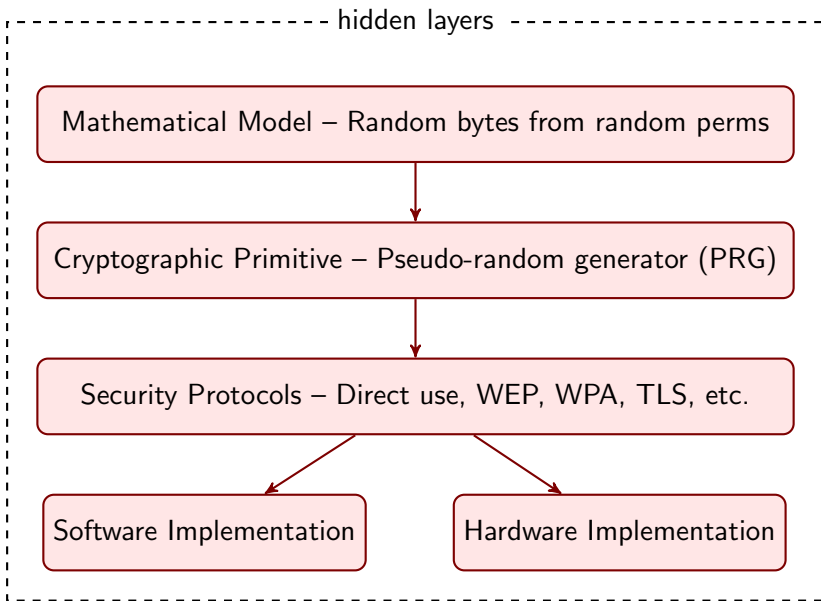
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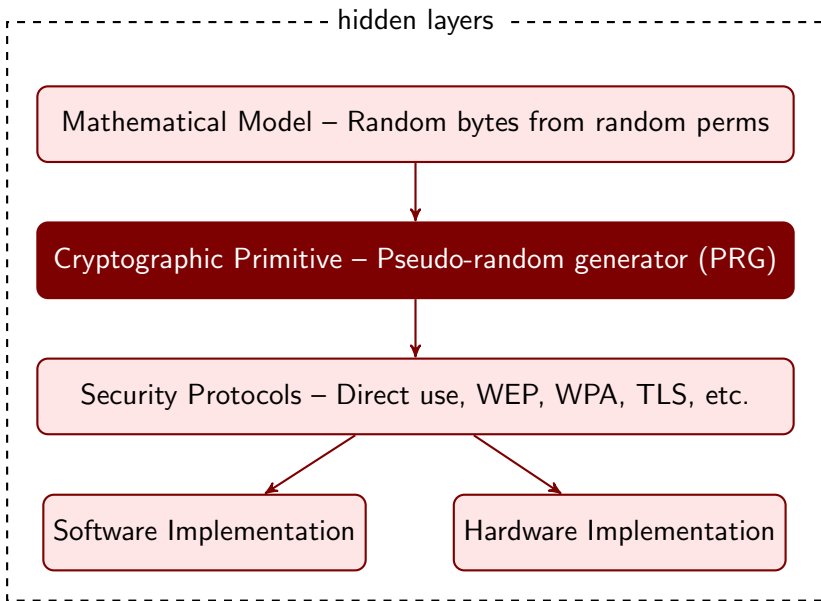
One must *cultivate* this cipher!

# Part I

## Analysis of RC4

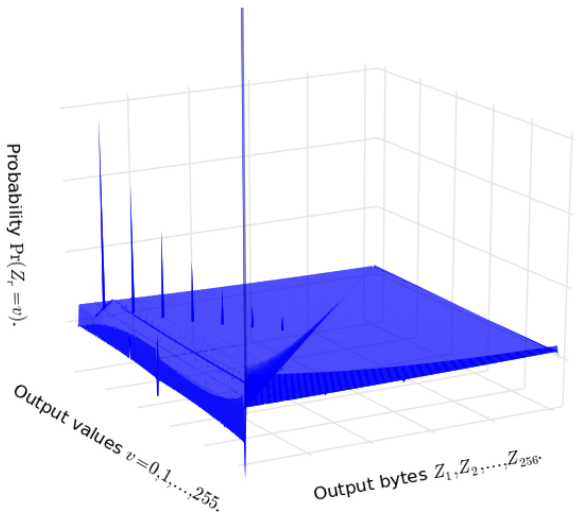






PRG output should be *indistinguishable* from truly random bitstream!

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## Broadcast attack on RC4

---

Encryption using RC4 is typically

$$E(k, P) : C \leftarrow P \oplus RC4(k)$$

$$C_1 = P_1 \oplus Z_1, \quad C_2 = P_2 \oplus Z_2, \quad C_3 = P_3 \oplus Z_3, \quad \dots$$

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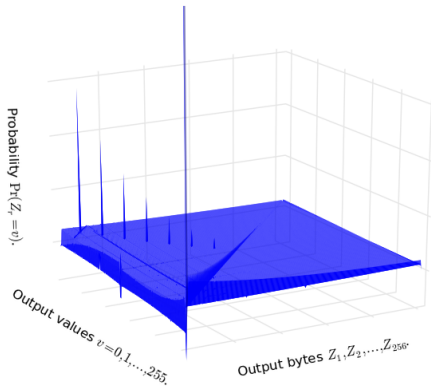
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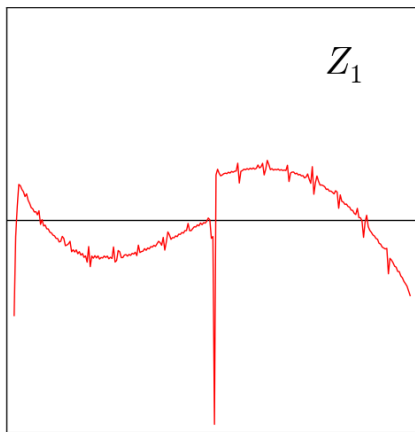
Plaintext recovery

- Gather multiple  $C$  and compute  $P_2 = \text{majority}\{C_2\}$
- Attack will be successful if number of  $C$  is in  $\Omega(N)$





## Non-randomness in initial bytes

Non-randomness in  $Z_1$ 

$$\Pr(Z_1 = v)$$

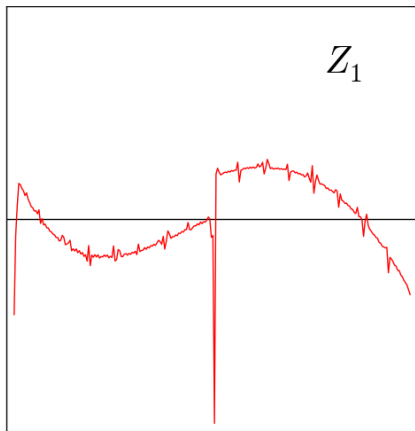
$$v = 0, 1, \dots, 255$$

Major biases

Sinusoidal distribution

$$\Pr(Z_1 = 0) \approx \frac{1}{N} - \frac{1}{N^2}$$

$$\Pr(Z_1 = 129) \approx \frac{1}{N} + \frac{2}{N^2}$$

Non-randomness in  $Z_1$ 

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Mironov, Crypto 2002

## Negative bias in ( $Z_1 = 0$ )

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### Theorem

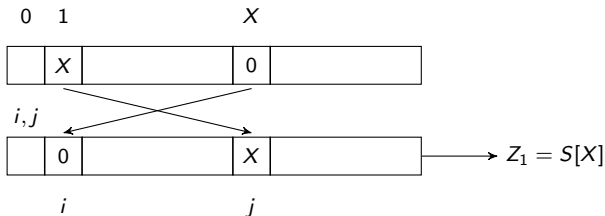
*Suppose the initial permutation of RC4 PRGA is a random permutation of  $\{0, 1, \dots, N - 1\}$ . Then  $\Pr(Z_1 = 0) \approx 1/N - 1/N^2$ .*

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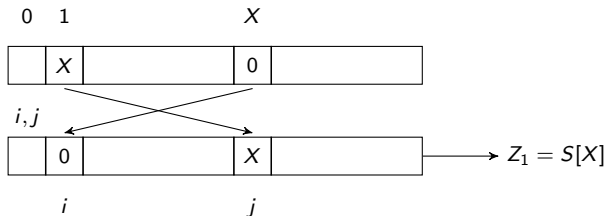


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$$\Pr(Z_1 = 0) \approx 0 \cdot 1/N + 1/N \cdot (1 - 1/N) = 1/N - 1/N^2$$

## Complete distribution of $Z_1$

---

### Theorem

For regular RC4, the probability distribution of  $Z_1$  is as follows,

$$\Pr(Z_1 = v) = Q_v + \sum_{X \in \mathcal{L}_v} \sum_{Y \in \mathcal{T}_{v,X}} \Pr(S_0[1] = X \wedge S_0[X] = Y \wedge S_0[X + Y] = v),$$

$$\text{with } Q_v = \begin{cases} \Pr(S_0[1] = 1 \wedge S_0[2] = 0), & \text{if } v = 0; \\ \Pr(S_0[1] = 0 \wedge S_0[0] = 1), & \text{if } v = 1; \\ \Pr(S_0[1] = 1 \wedge S_0[2] = v) \\ \quad + \Pr(S_0[1] = v \wedge S_0[v] = 0) \\ \quad + \Pr(S_0[1] = 1 - v \wedge S_0[1 - v] = v), & \text{otherwise.} \end{cases}$$

where  $v \in \{0, \dots, N - 1\}$ ,  $\mathcal{L}_v = \{0, 1, \dots, N - 1\} \setminus \{1, v\}$ ,

$\mathcal{T}_{v,X} = \{0, 1, \dots, N - 1\} \setminus \{0, X, 1 - X, v\}$ .

## Complete distribution of $Z_1$

---

Idea for the proof.

One may write

$$\begin{aligned} Z_1 &= S_1[S_1[i_1] + S_1[j_1]] = S_1[S_0[j_1] + S_0[i_1]] \\ &= S_1[S_0[S_0[1]] + S_0[1]] = S_1[Y + X], \text{ where } X = S_0[1], Y = S_0[X] \end{aligned}$$



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and thus compute

$$\Pr(Z_1 = v) = \sum_{X=0}^{N-1} \sum_{Y=0}^{N-1} \Pr(S_0[1] = X \wedge S_0[X] = Y \wedge S_1[X+Y] = v).$$

## Complete distribution of $Z_1$

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We have a known distribution for  $S_0[u] = v$  (Mantin, 2001).

Thus the goal is to reduce the term  $S_1[X+Y]$  to the state  $S_0$ .

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Note that

- $S_1$  is different from  $S_0$  in at most two places,  $i_1 = 1$  and  $j_1 = X$ .
- Special cases for  $X + Y = 1$  and  $X + Y = X$  must be considered.

## Complete distribution of $Z_1$

---

Idea for the proof.

Special cases depending on  $X, Y$

- $X + Y = 1$  if and only if  $Y = 1 - X$ , which implies

$$Z_1 = S_1[1] = S_1[i_1] = S_0[j_1] = S_0[X] = Y = 1 - X$$

- $X + Y = X$  if and only if  $Y = 0$ , which implies

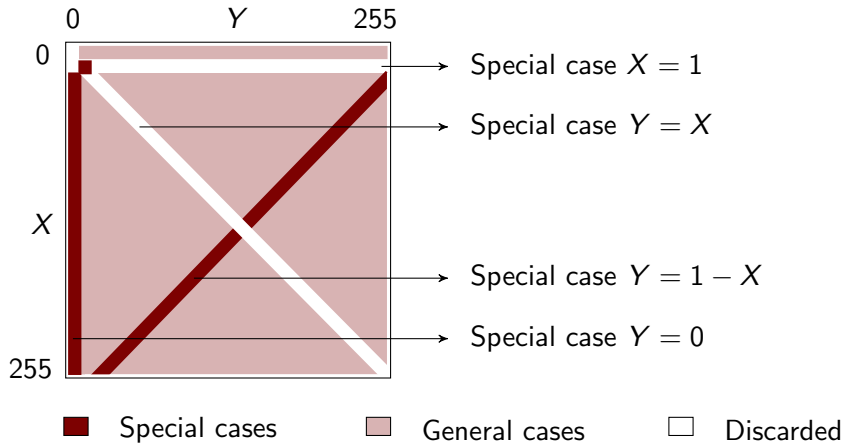
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- $X = 1$  if and only if  $Y = X$ , which implies

$$Z_1 = S_1[X + Y] = S_0[X + Y] = S_0[1 + 1] = S_0[2]$$

Complete distribution of  $Z_1$ 

Idea for the proof.



## Complete distribution of $Z_1$

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Idea for the proof.

$$\begin{aligned}
 \Pr(Z_1 = v) &= \sum_{X=0}^{N-1} \Pr(S_0[1] = X \wedge S_0[X] = 1 - X \wedge 1 - X = v) \\
 &+ \sum_{X=0}^{N-1} \Pr(S_0[1] = X \wedge S_0[X] = 0 \wedge X = v) \\
 &+ \Pr(S_0[1] = 1 \wedge S_0[2] = v) \\
 &+ \sum_{X \neq 1} \sum_{Y \neq 0, X, 1-X} \Pr(S_0[1] = X \wedge S_0[X] = Y \wedge S_0[X + Y] = v).
 \end{aligned}$$

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 \end{aligned}$$

The first summation term reduces to a single point ( $X = 1 - v, Y = v$ ), as we fix  $1 - X = v$  and  $Y = 1 - X$ .

## Complete distribution of $Z_1$

---

Idea for the proof.

$$\begin{aligned}
 \Pr(Z_1 = v) &= \Pr(S_0[1] = 1 - v \wedge S_0[1 - v] = v) \\
 &+ \sum_{X=0}^{N-1} \Pr(S_0[1] = X \wedge S_0[X] = 0 \wedge X = v) \\
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The second summation, similarly, reduces to point ( $X = v, Y = 0$ ).



## Complete distribution of $Z_1$

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Idea for the proof.

$$\begin{aligned} \Pr(Z_1 = v) &= \Pr(S_0[1] = 1 - v \wedge S_0[1 - v] = v) \\ &\quad + \Pr(S_0[1] = v \wedge S_0[v] = 0) \\ &\quad + \Pr(S_0[1] = 1 \wedge S_0[2] = v) \\ &\quad + \sum_{X \neq 1} \sum_{Y \neq 0, X, 1-X} \Pr(S_0[1] = X \wedge S_0[X] = Y \wedge S_0[X + Y] = v). \end{aligned}$$

Finally, we get two impossible conditions on the double summation:  
 $(X = v, Y \neq 0) \Rightarrow (Z_1 \neq v)$  and  $(X \neq 1 - v, Y = v) \Rightarrow (Z_1 \neq v)$ .

## Complete distribution of $Z_1$

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- $v = 1$  reduces the first three terms to  $\Pr(S_0[1] = 0 \wedge S_0[0] = 1)$ .

## Complete distribution of $Z_1$

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- $v = 1$  reduces the first three terms to  $\Pr(S_0[1] = 0 \wedge S_0[0] = 1)$ .
- $v \neq 0, 1$  keeps all the first three terms intact.

## Complete distribution of $Z_1$

---

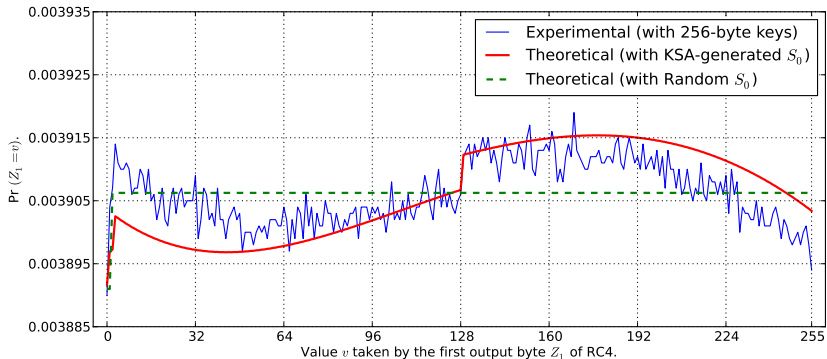
Hence the final expression

$$\Pr(Z_1 = v) = Q_v + \sum_{X \in \mathcal{L}_v} \sum_{Y \in \mathcal{T}_{v,X}} \Pr(S_0[1] = X \wedge S_0[X] = Y \wedge S_0[X + Y] = v),$$

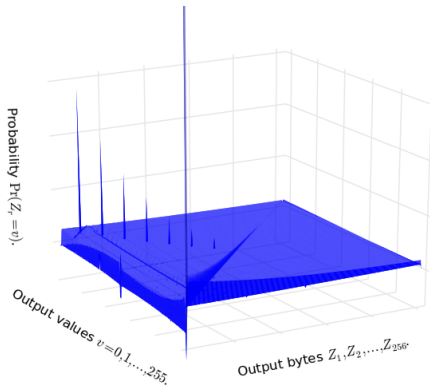
$$\text{with } Q_v = \begin{cases} \Pr(S_0[1] = 1 \wedge S_0[2] = 0), & \text{if } v = 0; \\ \Pr(S_0[1] = 0 \wedge S_0[0] = 1), & \text{if } v = 1; \\ \Pr(S_0[1] = 1 \wedge S_0[2] = v) \\ \quad + \Pr(S_0[1] = v \wedge S_0[v] = 0) \\ \quad + \Pr(S_0[1] = 1 - v \wedge S_0[1 - v] = v), & \text{otherwise.} \end{cases}$$

where  $v \in \{0, \dots, N - 1\}$ ,  $\mathcal{L}_v = \{0, 1, \dots, N - 1\} \setminus \{1, v\}$ ,

$\mathcal{T}_{v,X} = \{0, 1, \dots, N - 1\} \setminus \{0, X, 1 - X, v\}$ .

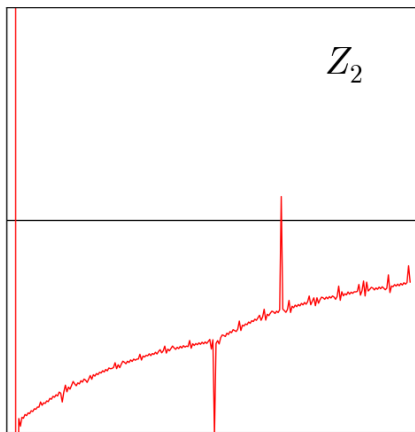
Complete distribution of  $Z_1$ 

Observed by Mironov in 2002. Proved by SMPS in 2013.



## Other initial bytes of RC4



Non-randomness in  $Z_2$ 

$$\Pr(Z_2 = v)$$

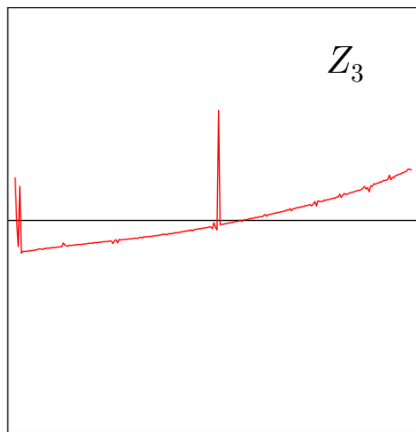
$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_2 = 0) \approx \frac{2}{N}$$

$$\Pr(Z_2 = 129) \approx \frac{1}{N} - \frac{2}{N^2}$$

$$\Pr(Z_2 = 172) \approx \frac{1}{N} + \frac{0.2}{N^2}$$

Non-randomness in  $Z_3$ 

$$\Pr(Z_3 = v)$$

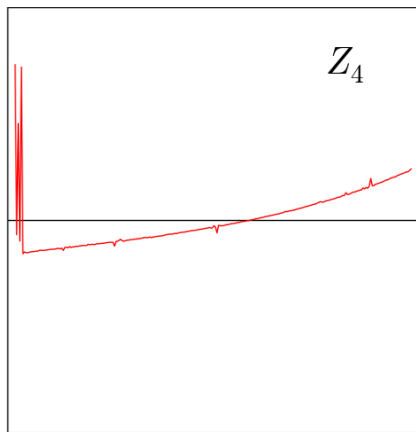
$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_3 = 0) \approx \frac{1}{N} + \frac{0.3}{N^2}$$

$$\Pr(Z_3 = 3) \approx \frac{1}{N} + \frac{0.3}{N^2}$$

$$\Pr(Z_3 = 131) \approx \frac{1}{N} + \frac{2}{N^2}$$

Non-randomness in  $Z_4$ 

$$\Pr(Z_4 = v)$$

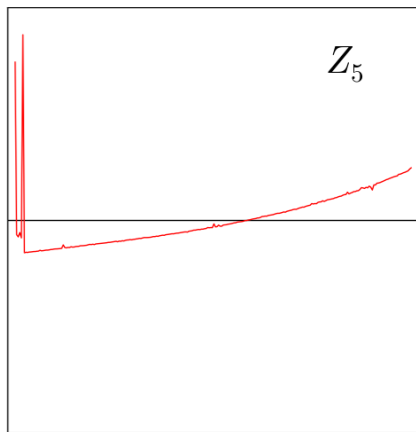
$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_4 = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_4 = 4) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_4 = 2) \approx \frac{1}{N} + \frac{0.8}{N^2}$$

Non-randomness in  $Z_5$ 

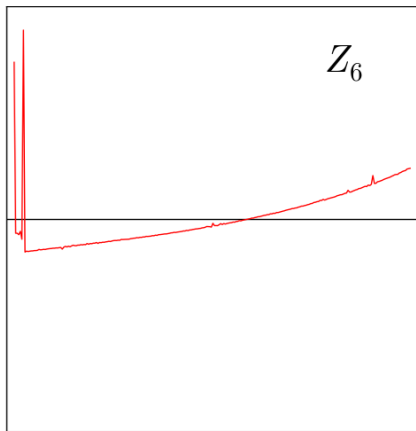
$$\Pr(Z_5 = v)$$

$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_5 = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_5 = 5) \approx \frac{1}{N} + \frac{1}{N^2}$$

Non-randomness in  $Z_6$ 

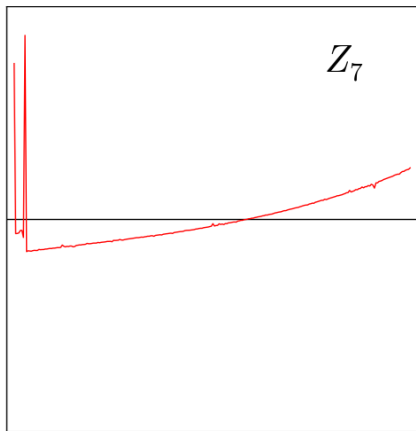
$$\Pr(Z_6 = v)$$

$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_6 = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_6 = 6) \approx \frac{1}{N} + \frac{1}{N^2}$$

Non-randomness in  $Z_7$ 

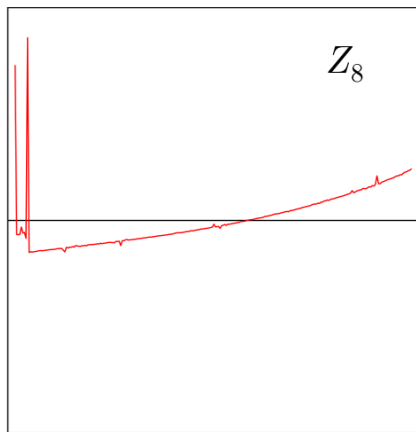
$$\Pr(Z_7 = v)$$

$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_7 = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_7 = 7) \approx \frac{1}{N} + \frac{1}{N^2}$$

Non-randomness in  $Z_8$ 

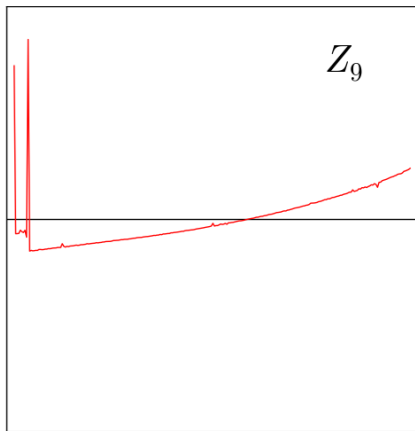
$$\Pr(Z_8 = v)$$

$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_8 = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_8 = 8) \approx \frac{1}{N} + \frac{1}{N^2}$$

Non-randomness in  $Z_9$ 

$$\Pr(Z_9 = v)$$

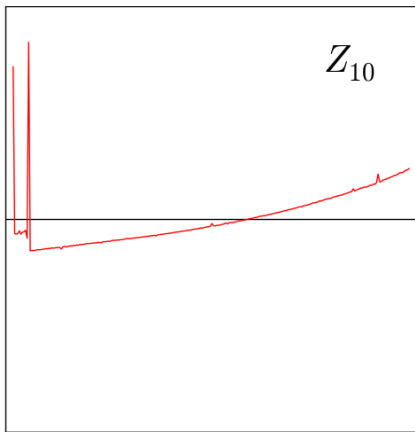
$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_9 = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_9 = 9) \approx \frac{1}{N} + \frac{1}{N^2}$$



Non-randomness in  $Z_{10}$ 

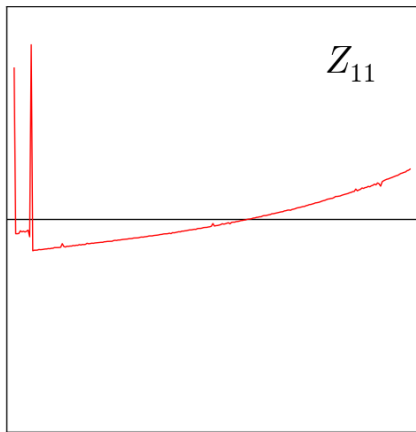
$$\Pr(Z_{10} = v)$$

$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_{10} = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_{10} = 10) \approx \frac{1}{N} + \frac{1}{N^2}$$

Non-randomness in  $Z_{11}$ 

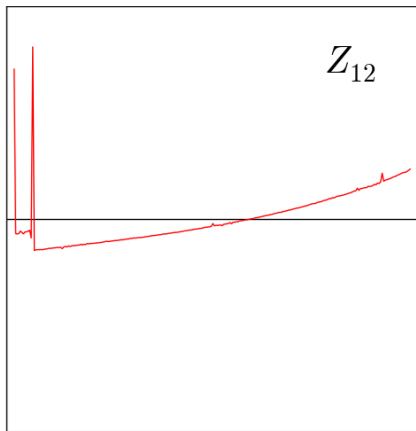
$$\Pr(Z_{11} = v)$$

$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_{11} = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_{11} = 11) \approx \frac{1}{N} + \frac{1}{N^2}$$

Non-randomness in  $Z_{12}$ 

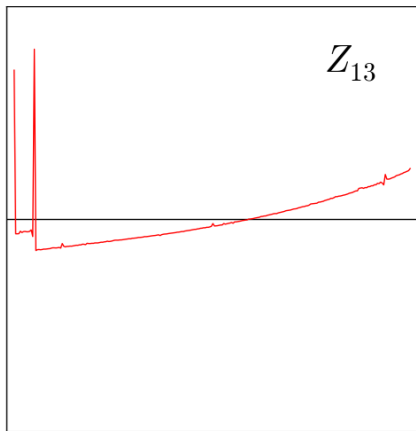
$$\Pr(Z_{12} = v)$$

$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_{12} = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_{12} = 12) \approx \frac{1}{N} + \frac{1}{N^2}$$

Non-randomness in  $Z_{13}$ 

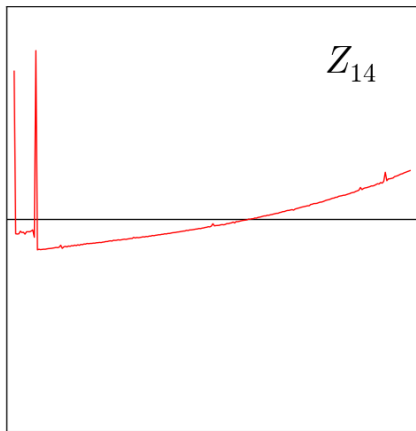
$$\Pr(Z_{13} = v)$$

$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_{13} = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_{13} = 13) \approx \frac{1}{N} + \frac{1}{N^2}$$

Non-randomness in  $Z_{14}$ 

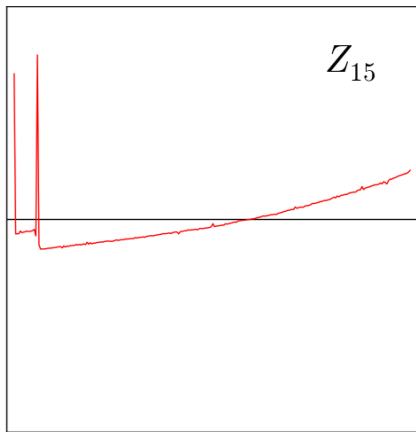
$$\Pr(Z_{14} = v)$$

$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_{14} = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_{14} = 14) \approx \frac{1}{N} + \frac{1}{N^2}$$

Non-randomness in  $Z_{15}$ 

$$\Pr(Z_{15} = v)$$

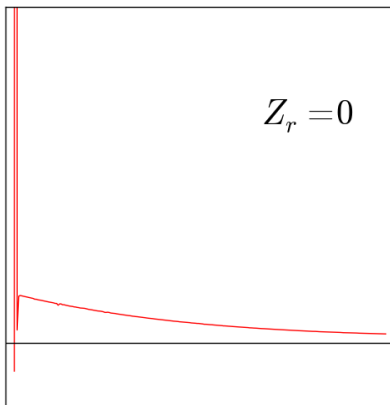
$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_{15} = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

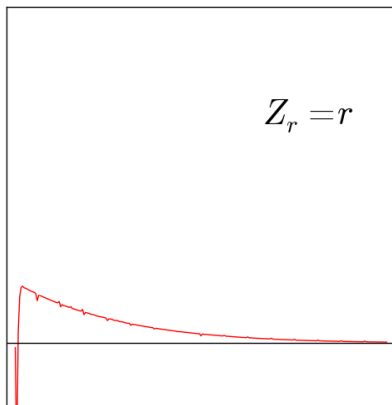
$$\Pr(Z_{15} = 15) \approx \frac{1}{N} + \frac{1}{N^2}$$

## Non-randomness in initial bytes



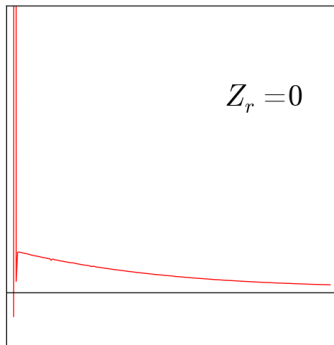
$$\Pr(Z_r = 0)$$

$$r = 1, 2, \dots, 255$$



$$\Pr(Z_r = r)$$

$$r = 1, 2, \dots, 255$$



## Zero-bias of initial bytes



## Zero-bias of initial bytes

---

- Mantin-Shamir discovered and proved the ( $Z_2 = 0$ ) bias in 2001.
- They claimed there are no biases towards zero for bytes 3 to 255.
- We revisit their work and contradict this claim in 2011.

## Zero-bias of initial bytes

---

- Mantin-Shamir discovered and proved the ( $Z_2 = 0$ ) bias in 2001.
- They claimed there are no biases towards zero for bytes 3 to 255.
- We revisit their work and contradict this claim in 2011.

### Theorem

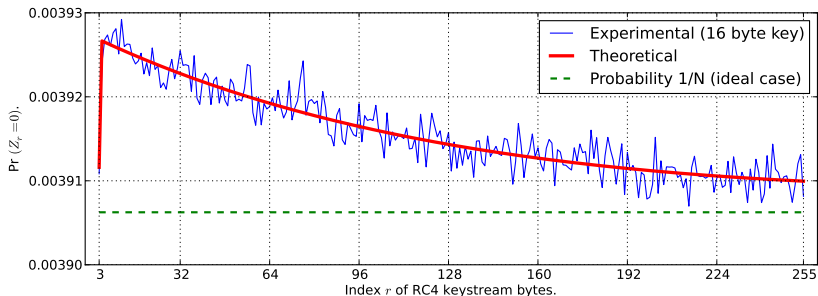
In PRGA rounds  $3 \leq r \leq N - 1$ , probability  $\Pr(Z_r = 0)$  is:

$$\Pr(Z_r = 0) \approx \frac{1}{N} + \frac{c_r}{N^2},$$

$$\text{where } c_r = \begin{cases} \frac{N}{N-1} (N \cdot \Pr(S_{r-1}[r] = r) - 1) - \frac{N-2}{N-1}, & \text{for } r = 3; \\ \frac{N}{N-1} (N \cdot \Pr(S_{r-1}[r] = r) - 1), & \text{otherwise.} \end{cases}$$

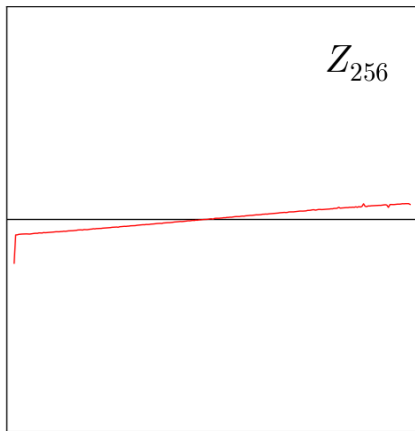
## Zero-bias of initial bytes

- Mantin-Shamir discovered and proved the ( $Z_2 = 0$ ) bias in 2001.
- They claimed there are no biases towards zero for bytes 3 to 255.
- We revisit their work and contradict this claim in 2011.



Zero-bias after byte 255

---

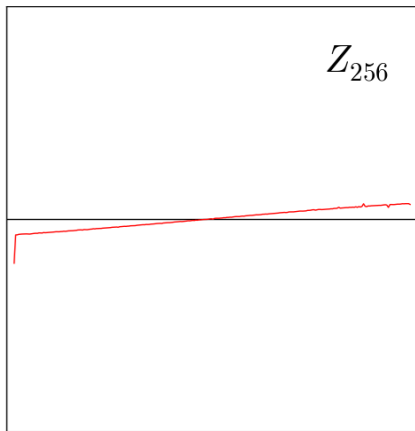


$$\Pr(Z_{256} = v)$$
$$v = 0, 1, \dots, 255$$

We proved

$$\Pr(Z_{256} = 0) \approx \frac{1}{N} - \frac{0.4}{N^2}$$

## Zero-bias after byte 255



$$\Pr(Z_{256} = v)$$

$$v = 0, 1, \dots, 255$$

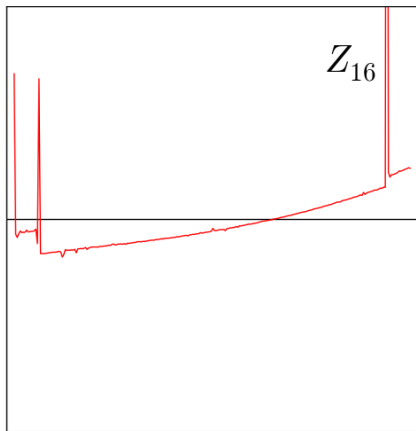
We proved

$$\Pr(Z_{256} = 0) \approx \frac{1}{N} - \frac{0.4}{N^2}$$

We also proved

$$\Pr(Z_{257} = 0) \approx \frac{1}{N} + \frac{0.35}{N^2}$$

**Something weird happens  
at the 16-th byte**

Strange bias in ( $Z_{16} = 240$ )

$$\Pr(Z_{16} = v)$$

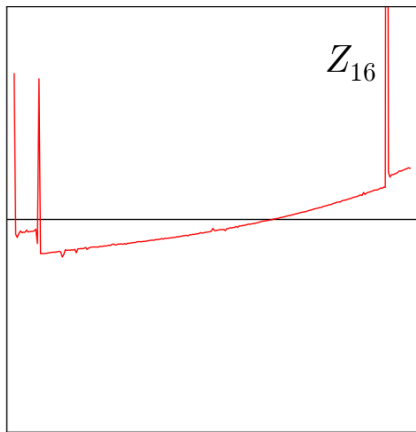
$$v = 0, 1, \dots, 255$$

Major biases

$$\Pr(Z_{16} = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_{16} = 16) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_{16} = 240) \approx \frac{1}{N} + \frac{9}{N^2}$$

Strange bias in ( $Z_{16} = 240$ )

$$\Pr(Z_{16} = v)$$

$$v = 0, 1, \dots, 255$$

Major biases

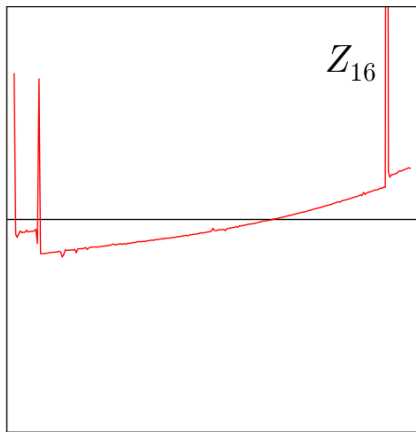
$$\Pr(Z_{16} = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_{16} = 16) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_{16} = 240) \approx \frac{1}{N} + \frac{9}{N^2}$$

Why 16?



Strange bias in ( $Z_{16} = 240$ )

$$\Pr(Z_{16} = v)$$

$$v = 0, 1, \dots, 255$$

Major biases

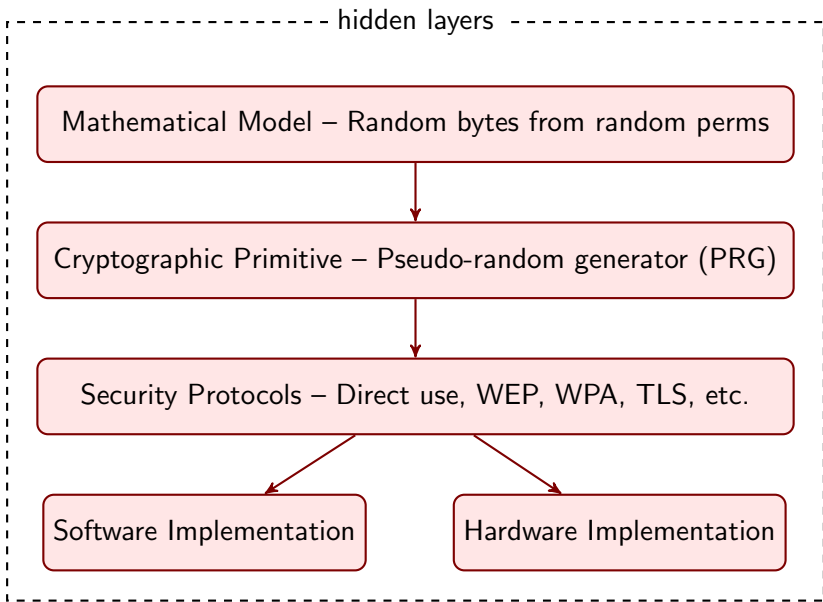
$$\Pr(Z_{16} = 0) \approx \frac{1}{N} + \frac{1}{N^2}$$

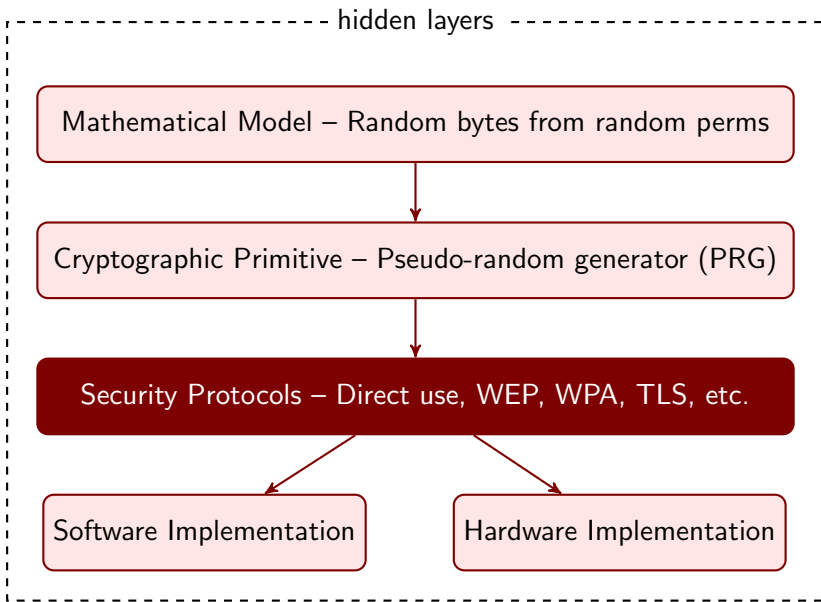
$$\Pr(Z_{16} = 16) \approx \frac{1}{N} + \frac{1}{N^2}$$

$$\Pr(Z_{16} = 240) \approx \frac{1}{N} + \frac{9}{N^2}$$

Why 16?

$240 \equiv -16$





# RC4 in Practice

## RC4 in practice

---

For the KSA step  $j = j + S[i] + K[i]$ , we require 256-byte  $K$  array. However in practice, the most typical key-size for RC4 is 128 bits.

## RC4 in practice

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 However in practice, the most typical key-size for RC4 is 128 bits.

KEY EXPANSION:  $K[i] = \text{RC4KEY}[i \bmod l]$  for  $i = 0, 1, 2, \dots, 255$ ,  
 where  $l$  is the length (in bytes) of the secret key



## RC4 in practice

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For the KSA step  $j = j + S[i] + K[i]$ , we require 256-byte  $K$  array. However in practice, the most typical key-size for RC4 is 128 bits.

KEY EXPANSION:  $K[i] = \text{RC4KEY}[i \bmod l]$  for  $i = 0, 1, 2, \dots, 255$ , where  $l$  is the length (in bytes) of the secret key



Typical length of the secret key:  $l = 128 \text{ bits} = 16 \text{ bytes}$

Intuition: This keylength of  $l = 16$  may have reflected in the  $Z_{16}$  bias.

# Discovery and proof of keylength-dependent biases



## Keylength-dependent distinguisher of RC4

---

$\Pr(Z_l = -l) > \frac{1}{N} + \frac{1}{N^2}$  for all practical keylengths  $l = 5, 6, \dots, 30$ .

## Keylength-dependent distinguisher of RC4

---

$\Pr(Z_l = -l) > \frac{1}{N} + \frac{1}{N^2}$  for all practical keylengths  $l = 5, 6, \dots, 30$ .

### Theorem

Suppose that  $l$  is the length of the secret key of RC4. Then

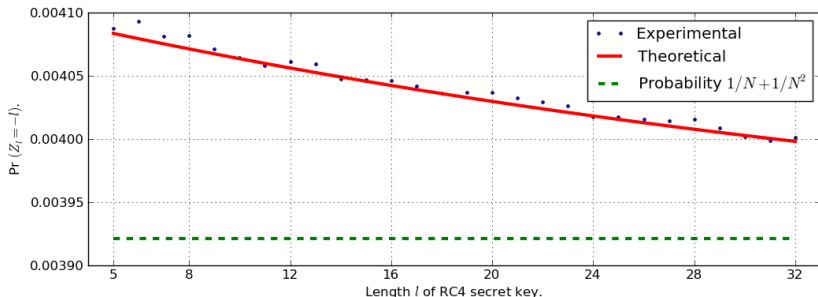
$$\Pr(Z_l = -l) \approx \frac{1}{N^2} + \left(1 - \frac{1}{N^2}\right) \gamma_l + (1 - \delta_l) \frac{1}{N},$$

where  $\gamma_l = \frac{1}{N^2} \left(1 - \frac{l+1}{N}\right) \sum_{x=l+1}^{N-1} \left(1 - \frac{1}{N}\right)^x \left(1 - \frac{2}{N}\right)^{x-l} \left(1 - \frac{3}{N}\right)^{N-x+2l-4}$  and

$$\delta_l = \Pr(S_1[l] = 0) \left(1 - \frac{1}{N}\right)^{l-2} + \sum_{t=2}^{l-1} \sum_{w=0}^{l-t} \frac{\Pr(S_1[t]=0)}{w! \cdot N} \left(\frac{l-t-1}{N}\right)^w \left(1 - \frac{1}{N}\right)^{l-3-w}.$$

## Keylength-dependent distinguisher of RC4

$\Pr(Z_l = -l) > \frac{1}{N} + \frac{1}{N^2}$  for all practical keylengths  $l = 5, 6, \dots, 32$ .



## Extended keylength-dependent biases

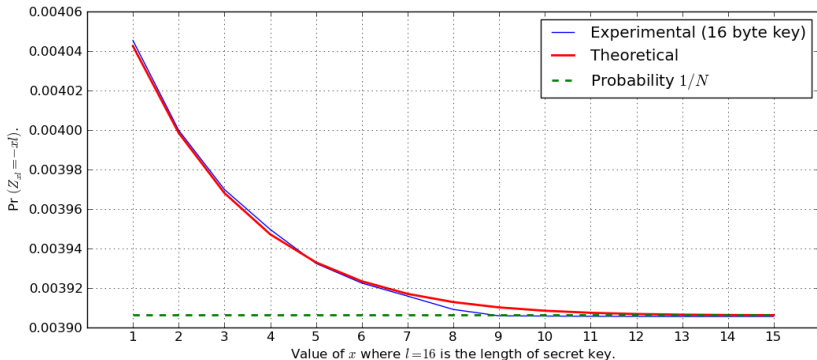
---

$\Pr(Z_{x/l} = -x/l) > \frac{1}{N}$  for  $l = 5, 6, \dots, 32$  and  $x = 1, 2, \dots, \lfloor \frac{N}{l} \rfloor$

## Extended keylength-dependent biases

$$\Pr(Z_{x/l} = -x/l) > \frac{1}{N} \text{ for } l = 5, 6, \dots, 32 \text{ and } x = 1, 2, \dots, \lfloor \frac{N}{l} \rfloor$$

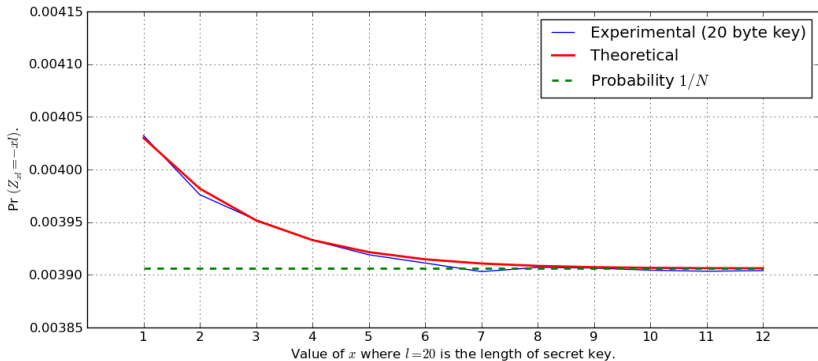
Example for  $l = 16$



## Extended keylength-dependent biases

$\Pr(Z_{x/l} = -x/l) > \frac{1}{N}$  for  $l = 5, 6, \dots, 32$  and  $x = 1, 2, \dots, \lfloor \frac{N}{l} \rfloor$

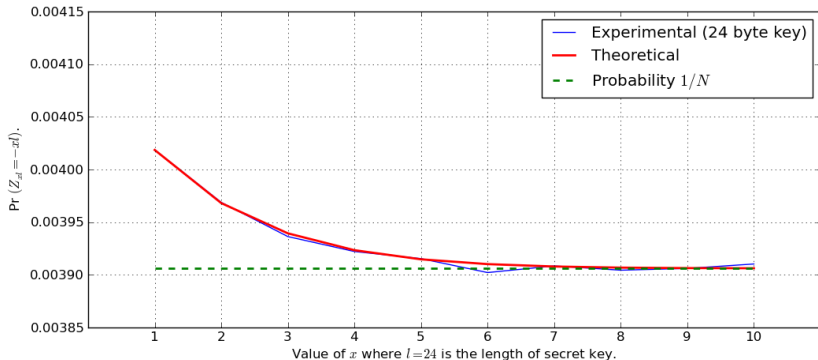
Example for  $l = 20$



## Extended keylength-dependent biases

$\Pr(Z_{x/l} = -x/l) > \frac{1}{N}$  for  $l = 5, 6, \dots, 32$  and  $x = 1, 2, \dots, \lfloor \frac{N}{l} \rfloor$

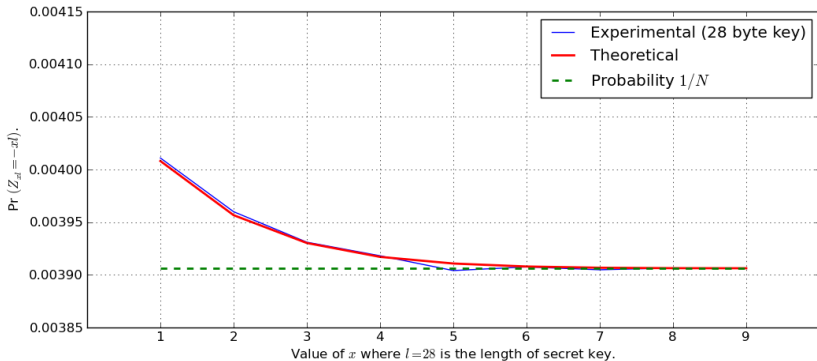
Example for  $l = 24$



## Extended keylength-dependent biases

$\Pr(Z_{x/l} = -x/l) > \frac{1}{N}$  for  $l = 5, 6, \dots, 32$  and  $x = 1, 2, \dots, \lfloor \frac{N}{l} \rfloor$

Example for  $l = 28$

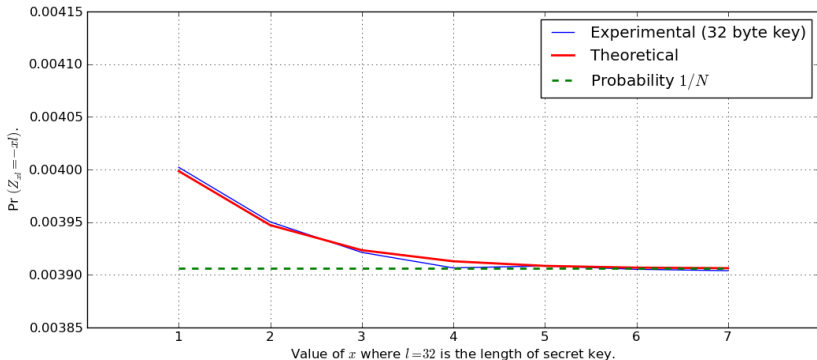


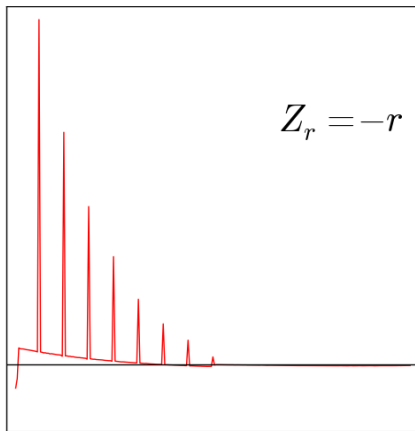


## Extended keylength-dependent biases

$\Pr(Z_{x/l} = -x/l) > \frac{1}{N}$  for  $l = 5, 6, \dots, 32$  and  $x = 1, 2, \dots, \lfloor \frac{N}{l} \rfloor$

Example for  $l = 32$



Keylength-dependent biases for  $l = 16$ 

$$\Pr(Z_r = -r)$$

$$r = 1, \dots, 255$$

Major biases

$$\Pr(Z_{16} = 240) \approx \frac{1}{N} + \frac{9}{N^2}$$

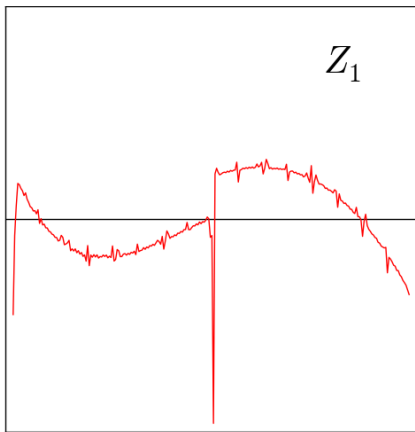
$$\Pr(Z_{32} = 224) \approx \frac{1}{N} + \frac{6}{N^2}$$

$$\Pr(Z_{48} = 208) \approx \frac{1}{N} + \frac{4}{N^2}$$

$$\Pr(Z_{64} = 192) \approx \frac{1}{N} + \frac{3}{N^2}$$

$$\Pr(Z_{80} = 176) \approx \frac{1}{N} + \frac{2}{N^2}$$

**Keylength affects  $Z_1$  too**

Keylength-dependence in  $Z_1$ 

$$\Pr(Z_1 = v)$$

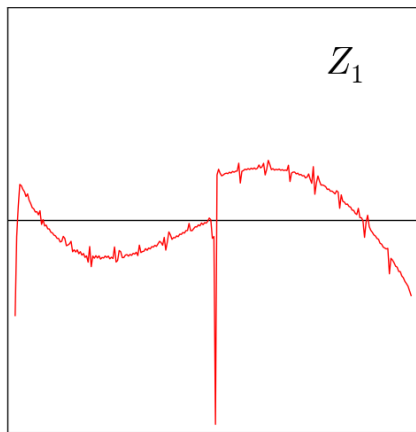
$$v = 0, 1, \dots, 255$$

Major biases

Sinusoidal distribution

$$\Pr(Z_1 = 0) \approx \frac{1}{N} - \frac{1}{N^2}$$

$$\Pr(Z_1 = 129) \approx \frac{1}{N} - \frac{2}{N^2}$$

Keylength-dependence in  $Z_1$ 

$$\Pr(Z_1 = v)$$

$$v = 0, 1, \dots, 255$$

Major biases

Sinusoidal distribution

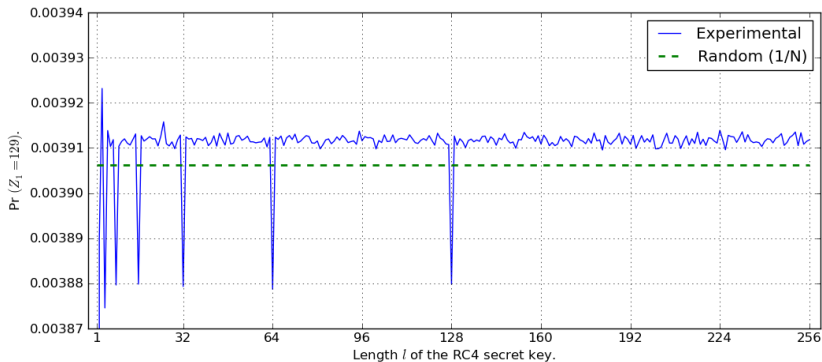
$$\Pr(Z_1 = 0) \approx \frac{1}{N} - \frac{1}{N^2}$$

$$\Pr(Z_1 = 129) \approx \frac{1}{N} + \frac{2}{N^2}$$

For  $l = 16$ , not for  $l = 256$

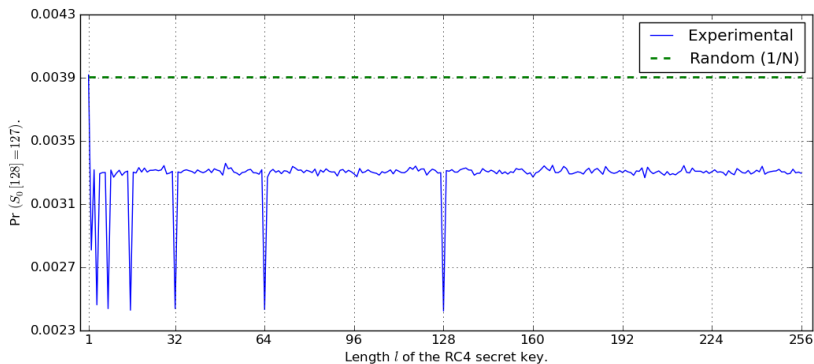
## Keylength-dependence in $Z_1$

Bias at ( $Z_1 = 129$ ) is present only for  $l = 2, 4, 8, 16, 32, 64, 128$



## Keylength-dependence in $S_0$

Bias at ( $S_0[128] = 127$ ) is present only for  $l = 2, 4, 8, 16, 32, 64, 128$



## Keylength-dependence in $S_0$

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( $S_0[128] = 127$ ) bias for  $l = 16$  was known as an *anomaly* since 2001.  
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### Theorem

*In practical RC4 with  $N = 256$ ,*

$$\Pr(S_0[128] = 127) \approx 0.63/N,$$

*if and only if  $l$  is a non-trivial factor of  $N = 256$ .*

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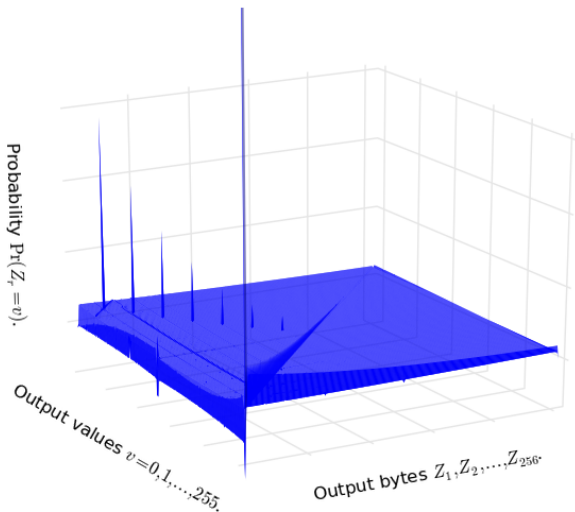
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*if and only if  $l$  is a non-trivial factor of  $N = 256$ .*

Intuition for the proof: The calculation for  $\Pr(S_0[128] = 127)$  behaves differently if  $K[128] = K[0]$  after key expansion; this happens with certainty if and only if  $l = 2, 4, 8, 16, 32, 64, 128$ .

# Practical implication of initial-byte biases

RC4 becomes weak against broadcast attack on initial plaintext bytes!



## Recent plaintext-recovery attacks

---

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AlFardan et al., 2013

- Recovery of all initial bytes using all initial byte biases.
- Broadcast attack on TLS using the same technique.



**Discard all problematic  
initial output bytes!**

## Long-term bias in RC4

---

Golic proved a bitwise correlation between  $Z_r$  and  $Z_{r+2}$  in 1997.

We prove a new periodic bitwise correlation between  $Z_r$  and  $Z_{r+2}$ .

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### Theorem

*Suppose that the permutation  $S_{wN}$  is truly random, then for  $w > 0$ ,*

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This is the first long-term byte-wise correlation (periodic) to be observed between two non-consecutive bytes.

# Biases related to the state-variables

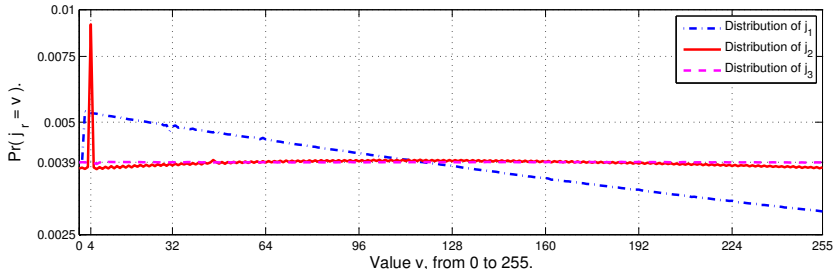
## State-dependent biases

Observed by SVV in 2010, proved by SMPS in 2011.

Type of Bias	Label by SVV'10	Biases proved
Specific Initial Rounds	"New_004"	$j_2 + S_2[j_2] = S_2[i_2] + Z_2$
	"New_noz_007"	$j_2 + S_2[j_2] = 6$
	"New_noz_009"	$j_2 + S_2[j_2] = S_2[i_2]$
	"New_noz_014"	$j_1 + S_1[i_1] = 2$
All Rounds ( $r$ -independent)	"New_noz_001"	$j_r + S_r[i_r] = i_r + S_r[j_r]$
	"New_noz_002"	$j_r + S_r[j_r] = i_r + S_r[i_r]$
All Initial Rounds ( $r$ -dependent)	"New_000"	$S_r[t_r] = t_r$
	"New_noz_004"	$S_r[i_r] = j_r$
	"New_noz_006"	$S_r[j_r] = i_r$

## Non-randomness of index $j$

We characterized the non-randomness in index  $j$  and in the process, discovered a new bias in ( $j_2 = 4$ ).



Index  $j$  behaves random from onwards  $j_3$ .

## Glimpse in RC4

---

We exploited the bias in  $(j_2 = 4)$  to get a short-term glimpse.

$$\Pr(S_2[2] = 4 - Z_2) \approx \frac{1}{N} + \frac{4/3}{N^2}.$$



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The best existing long-term glimpse was by Jenkins in 1996.

$$\Pr(S_r[j_r] = i_r - Z_r) = \Pr(S_r[i_r] = j_r - Z_r) \approx \frac{2}{N}$$

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We identified the proved a new long-term glimpse in 2013.

$$\Pr(S_r[r + 1] = N - 1 \mid Z_{r+1} = Z_r \wedge Z_{r+1} = r + 2) \approx \frac{3}{N}$$

# Contributions in RC4 Analysis

## Contributions in RC4 Analysis

---

Settling long-standing open problems	Ref.
1. Keylength dependent anomaly	Mantin, 2001
2. Long-term conditional glimpse	Jenkins, 1996
3. Distribution of $Z_1$	Mironov, 2002
4. Zero-bias of bytes $Z_3, \dots, Z_{255}$	MS, 2001
5. Long-term bias in non-consecutive bytes	Golic, 1997

## Contributions in RC4 Analysis

---

Providing theoretical validation of practical attacks	Ref.
1. Proving biases used in WEP and WPA attacks	SVV, 2010
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| Initiating new directions in RC4 analysis             | Ref.        |
| 1. Keylength-dependent biases in RC4                  | SMPS, 2013  |
| 2. Keylength-dependence in $Z_1$ bias                 | SSPM, 2013  |

# Part II

## Implementation of RC4

## Motivation for this Work

---

“In how many clocks a byte can be generated in RC4 PRGA?”

Most common approach

- 1 cycle for increment/computation of indices  $i, j$
- 1 cycle for swapping the values  $S[i]$  and  $S[j]$
- 1 cycle for reading the  $Z$  value from  $S$ -array

MOTIVATION: Can we get a better throughput?



## Design 1 – Loop unrolling

---

“One Byte per Clock throughput for RC4 PRGA”

3 → 1

- $N$  bytes of output in  $N + 2$  clock cycles
- Completion of RC4 KSA in 257 clock cycles
- Asymptotically *'one byte per clock cycle'*

## Design 1 – Loop unrolling

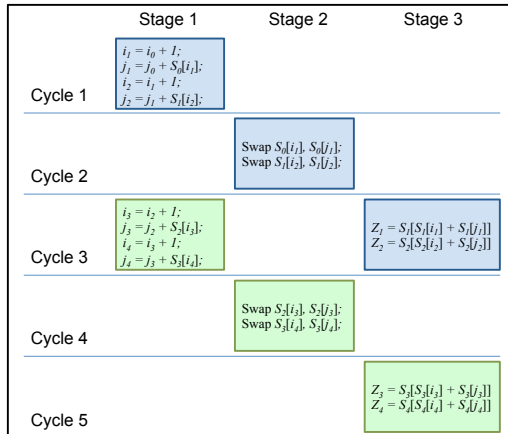
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“Combine two rounds of RC4 PRGA”

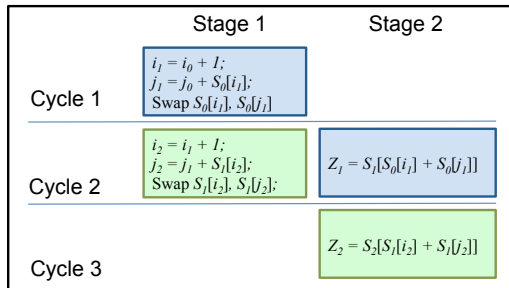
Steps	First Loop	Second Loop
1	$i_1 = i_0 + 1$	$i_2 = i_1 + 1 = i_0 + 2$
2	$j_1 = j_0 + S_0[i_1]$	$j_2 = j_1 + S_1[i_2] = j_0 + S_0[i_1] + S_1[i_2]$
3	Swap $S_0[i_1] \leftrightarrow S_0[j_1]$	Swap $S_1[i_2] \leftrightarrow S_1[j_2]$
4	$Z_1 = S_1[S_0[i_1] + S_0[j_1]]$	$Z_2 = S_2[S_1[i_2] + S_1[j_2]]$

- What if the indices overlap? (e.g.,  $j_1 = i_2$ )
- What about the ordering of *Swap* and *Output*?

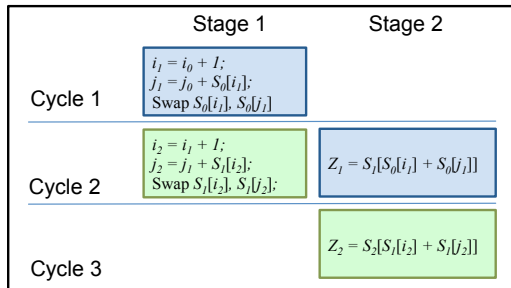
## Design 1 – Loop unrolling



## Design 1.5 – Simple hardware pipeline

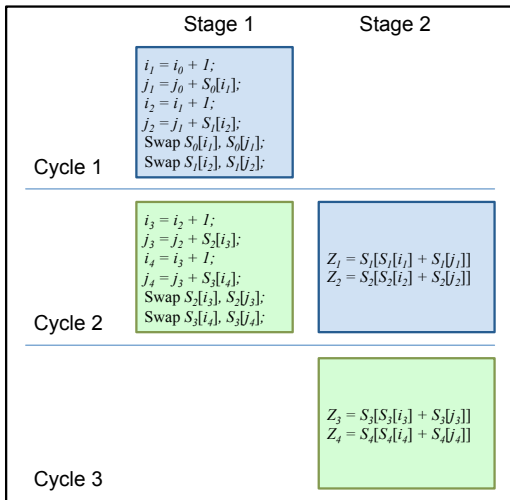


## Design 1.5 – Simple hardware pipeline



This approach is independent of the loop unrolling.  
 Is it possible to merge the two approaches?

## Design 2 – Hybrid approach



## Design 2 – Hybrid approach

---

“Two Bytes per Clock throughput for RC4 PRGA”

1  $\longrightarrow$  0.5

- $2N$  bytes of output in  $N + 1$  clock cycles
- Completion of RC4 KSA in 129 clock cycles
- Asymptotically *‘two bytes per clock cycle’*

## Contributions in RC4 Implementation

---

Improved the throughputs of common RC4 designs in literature.  
Matched the best throughput 1-byte-per-cycle from industry patents.  
Provided the best throughput 2-bytes-per-cycle design for RC4.

---

Year	Result in RC4 implementation	Ref.
2003	3 cycles-per-byte design based on custom pipeline	Kitsos
2003	3 cycles-per-byte design based on multi-port memory	Matthews
2008	1 cycle-per-byte design based on hardware pipelining	Matthews
2010	1 byte-per-cycle design based on loop unrolling	SSMS
2013	2 bytes-per-cycle design based on hardware pipelining combined with loop unrolling in a hybrid model	SCSMS

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# Open problems in RC4

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### Key collisions

- Theoretical construction of short colliding key-pairs.
- Search for collision with 16-byte key-pairs in RC4.

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### Key collisions

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### Key recovery

- Narrow the gap of theory and practice in terms of key recovery attacks on WEP and WPA.

### Anomaly pairs

- Characterization of all anomalies in RC4.
- Identify and prove all anomaly-dependent biases.

## Open problems in RC4

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- Analysis and improvement of existing results in state recovery.

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- Search for all significant biases of the form  $(Z_r \star Z_{r+x} = v)$ .

### Hardware implementation

- Area optimization by distributing  $S$ -array over memory banks.



# Publications

## Publications from the Thesis

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### RC4 ANALYSIS

1. Sourav Sen Gupta, Subhamoy Maitra, Goutam Paul, and Santanu Sarkar. (Non-)random sequences from (non-)random permutations – analysis of RC4 stream cipher. *Journal of Cryptology*, 2013.
2. Santanu Sarkar, Sourav Sen Gupta, Goutam Paul, and Subhamoy Maitra. Proving TLS-attack related open biases of RC4. *IACR ePrint*, 2013.
3. Subhamoy Maitra and Sourav Sen Gupta. New long-term glimpse of RC4 stream cipher. In *ICISS*. Springer LNCS, 2013.
4. Subhamoy Maitra, Goutam Paul, and Sourav Sen Gupta. Attack on broadcast RC4 revisited. In *FSE*. Springer LNCS, 2011.
5. Sourav Sen Gupta, Subhamoy Maitra, Goutam Paul, and Santanu Sarkar. Proof of empirical RC4 biases and new key correlations. In *Selected Areas in Cryptography*. Springer LNCS, 2011.

## Publications from the Thesis

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### RC4 IMPLEMENTATION

1. Sourav Sen Gupta, Anupam Chattopadhyay, Koushik Sinha, Subhamoy Maitra, and Bhabani P. Sinha. High-performance hardware implementation for RC4 stream cipher. *IEEE Transactions on Computers*, 2013.
2. Sourav Sen Gupta, Koushik Sinha, Subhamoy Maitra, and Bhabani P. Sinha. One byte per clock: A novel RC4 hardware. In *INDOCRYPT*. Springer LNCS, 2010.

Total: 2 journal papers, 4 conference papers, 1 ePrint report.

THANK YOU  
for your kind attention